Supplemental Appendix: Robustness Checks and Additional Results

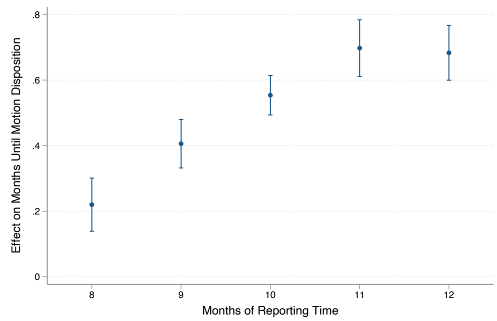

Figure A1. Non-parametric OLS Estimates

Note: Figure plots non-parametric OLS estimates of the effect of reporting time on months until motion disposition. Reporting time is defined as the number of months between the day on which a motion was filed and the earliest possible date on which it could appear on a six-month list. Estimates are obtained by regressing months until motion disposition on a series of reporting time dummies (one for each month of reporting time) in addition to motion and case controls. Error bars indicate 95% confidence intervals.

Table A1—: Balance Table: RD Estimates for Pre-Treatment Characteristics.

	Linear		Quadratic	
Dependent Var.	(1) No Controls	(2) Controls	(3) No Controls	(4) Controls
% Filed by Defendant	0.00 (0.01)	0.00 (0.01)	-0.00 (0.02)	-0.00 (0.02)
% Class Action	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)
Docket Entries Prior to Motion	-12.14 (11.56)	-14.39 (10.69)	-12.27 (12.86)	-14.51 (12.53)

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Table A1 – ...continued from previous page

	Linear		Quadratic	
	(1)	(2)	(3)	(4)
Dependent Var.	No Controls	Controls	No Controls	Controls
Days Prior to Motion Filing	-22.72	-18.72	-33.11	-35.94
	(18.26)	(14.49)	(25.56)	(24.52)
% Civil Rights Cases	0.01	0.01	-0.00	-0.00
	(0.01)	(0.01)	(0.01)	(0.01)
% Employment Cases	0.01*	0.01*	0.01	0.01
	(0.01)	(0.01)	(0.01)	(0.01)
% Prisoner Rights Cases	-0.02***	-0.02***	-0.02***	-0.02***
	(0.01)	(0.01)	(0.01)	(0.01)
% Intellectual Property Cases	-0.00	-0.00	-0.01*	-0.01*
	(0.00)	(0.00)	(0.00)	(0.00)
% Labor Cases	0.01**	0.01**	0.01**	0.01**
	(0.01)	(0.01)	(0.01)	(0.01)
% Real Property Cases	0.00	0.00	0.00	0.00
	(0.00)	(0.00)	(0.00)	(0.00)
% Soc. Sec. Appeals	-0.02	-0.01	-0.02	-0.01
	(0.01)	(0.01)	(0.01)	(0.02)
% Tort Cases	-0.01	-0.01	-0.01	-0.02
	(0.02)	(0.02)	(0.02)	(0.03)
% Contract Cases	0.01	0.01	0.01	0.01
	(0.01)	(0.01)	(0.01)	(0.01)
% Motion Filed by Pro Se Party	-0.01	-0.01	-0.01	-0.01
	(0.00)	(0.00)	(0.01)	(0.01)
% In Forma Pauperis	-0.01	-0.01	-0.01	-0.01
	(0.01)	(0.01)	(0.01)	(0.01)
% Diversity Jurisdiction	-0.00	-0.01	-0.02	-0.02

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Table A1 – ...continued from previous page

	Linear		Quadratic	
	(1)	(2)	(3)	(4)
Dependent Var.	No Controls	Controls	No Controls	Controls
	(0.02)	(0.02)	(0.03)	(0.03)
% Federal Question Jurisdiction	0.02*	0.02*	0.03*	0.03*
	(0.01)	(0.01)	(0.02)	(0.02)
% Government Defendant	-0.01	-0.01	-0.01	-0.01
	(0.01)	(0.01)	(0.02)	(0.02)
Assigned Judge Age (Years)	-0.33**	-0.35**	-0.39*	-0.41**
	(0.16)	(0.15)	(0.20)	(0.19)
% Assigned Female Judge	0.02	0.02*	0.03*	0.03**
	(0.01)	(0.01)	(0.02)	(0.02)
% Assigned Non-White Judge	0.01	0.01	0.01	0.01
	(0.01)	(0.01)	(0.01)	(0.01)
% Assigned Magistrate Judge	0.00	0.00	0.00	0.00
	(0.01)	(0.01)	(0.01)	(0.01)

Note: Table shows regression discontinuity estimates for various pre-treatment motion- and case-level characteristics. MSE-optimal bandwidths are selected using the procedure from Calonico et al. (2019) and based on a local linear regression. All columns estimated with first- or second-order local polynomials. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table A2—RD Estimates for Months Until Motion Disposition, Alternative Bandwidths

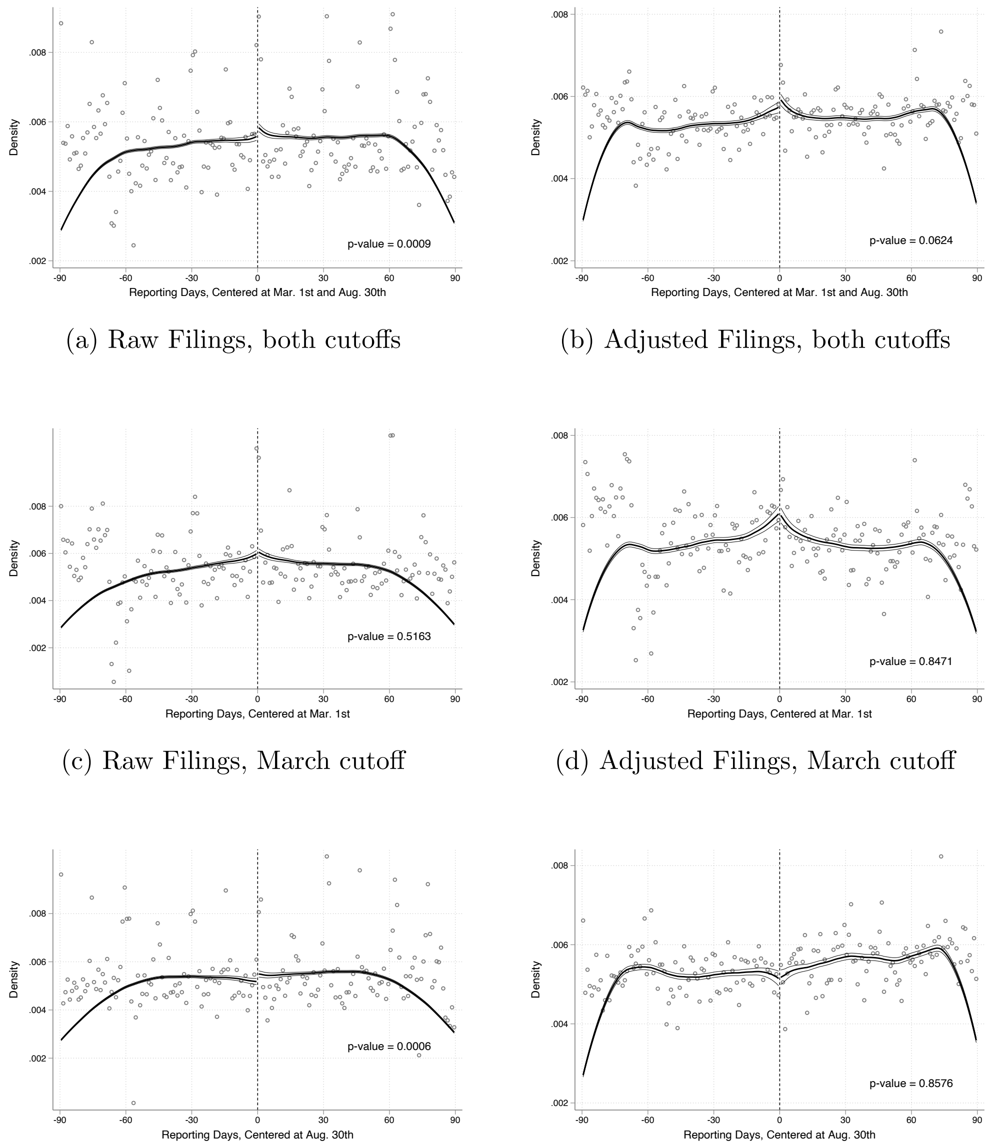
	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Filed After Cutoff	0.788***	0.723***	0.735***	0.746***	0.777***	0.792***	0.806***
	(0.115)	(0.113)	(0.108)	(0.100)	(0.091)	(0.084)	(0.078)
Dep. Var. mean	5.4	5.4	5.4	5.4	5.4	5.4	5.4
Polynomial order	First	First	First	First	First	First	First
Controls	✓	✓	✓	✓	✓	✓	✓
Bandwidth (days)	7.0	14.0	21.0	28.0	35.0	42.0	49.0
Effective N	15,958	31,318	48,173	63,142	82,440	96,850	113990
N (left)	7,264	14,963	23,190	30,654	40,223	47,563	55,998
N (right)	8,694	16,355	24,983	32,488	42,217	49,287	57,992

Note: Table presents RD estimates of the effect of reporting time on months until motion disposition using a variety alternative bandwidths. All columns are estimated with first first-order local polynomials and triangular kernels. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table A3—RD Estimates: Effect of six months additional reporting time on months until motion disposition, clustered at judge-level

	(1)	(2)	(3)	(4)
Filed After Cutoff	0.764***	0.785***	0.732***	0.758***
	(0.146)	(0.135)	(0.192)	(0.167)
Dep. Var. mean	5.4	5.4	5.4	5.4
Polynomial order	First	First	Second	Second
Controls		✓		✓
Bandwidth (days)	38.9	38.5	44.6	48.7
Effective N	90,754	90,549	103572	113813
N (left)	44,440	44,340	51,001	55,931
N (right)	46,314	46,209	52,571	57,882

Note: Table presents bias-corrected RD estimates of the effect of six months additional reporting time on months until motion disposition. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). Effects are estimated with first or second order local polynomials using triangular kernels. Columns indicating controls include list (i.e., March or September) fixed effects, a dummy for the part (plaintiff or defendant) filing the motion, and nature-of-suit and filing-year fixed effects. Robust standard errors clustered at the judge-level are in parentheses.



(a) Raw Filings, both cutoffs (b) Adjusted Filings, both cutoffs

(c) Raw Filings, March cutoff (d) Adjusted Filings, March cutoff

(e) Raw Filings, August cutoff (f) Adjusted Filings, August cutoff

Figure A2. Density of Motion Filing Dates Relative to Six-Month List Cut-offs

Note: Figures show the density of summary judgment motions filing dates relative to the six-month list maximum-reporting-time cut-offs of March 1 and August 30. Sample includes summary judgment motions filed from 2005–2014 with known dispositions. P-values displayed in the lower right are from McCrary density tests (McCrary, 2008). Panels (a)–(f) show raw and adjusted filing date densities for both, March, and August cutoffs respectively.

Table A4—Parametric RD Estimates: Effect of six months additional reporting time on months until motion disposition

	(1)	(2)	(3)	(4)
Filed After Cutoff	0.773***	0.814***	0.717***	0.766***
	(0.091)	(0.083)	(0.132)	(0.110)
Dep. Var. mean	5.32	5.32	5.33	5.33
Global poly. order	First	First	Second	Second
Controls		✓		✓
Bandwidth (days)	33.68	32.09	42.15	45.95
Effective N	80,264	78,120	99,208	106296
N (left)	39,125	38,144	48,741	52,285
N (right)	41,139	39,976	50,467	54,011

Note: Table presents bias-corrected RD estimates of the effect of six months additional reporting time on months until motion disposition. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014) but effects are estimated with first or second global polynomials. Columns indicating controls include list (i.e., March or September) fixed effects, a dummy for the part (plaintiff or defendant) filing the motion, and nature-of-suit and filing-year fixed effects. Robust standard errors clustered at the Julian date of filing are in parentheses.

Table A5—Donut RD Estimates for Months Until Motion Disposition

	(1)	(2)	(3)	(4)	(5)	(6)
Filed After Cutoff	0.778***	0.785***	0.790***	0.797***	0.802***	0.808***
	(0.091)	(0.088)	(0.085)	(0.082)	(0.079)	(0.077)
Donut radius (days)	3	6	9	12	15	18
Dep. Var. mean	5.4	5.4	5.4	5.4	5.4	5.4
Polynomial order	First	First	First	First	First	First
Controls	✓	✓	✓	✓	✓	✓
Bandwidth (days)	35.1	38.1	41.1	44.1	47.1	50.1
Effective N	84,507	90,684	96,850	103496	111594	118476
N (left)	41,251	44,393	47,563	50,947	54,724	58,180
N (right)	43,256	46,291	49,287	52,549	56,870	60,296

Note: Table presents "Donut" RD estimates of the effect of reporting time on months until motion disposition using a variety of donut radii. All columns are estimated with triangular kernels using first-order local polynomials.

Table A6—Placebo Tests: RD Estimates for Months Until Motion Disposition, Counterfactual Cutoffs

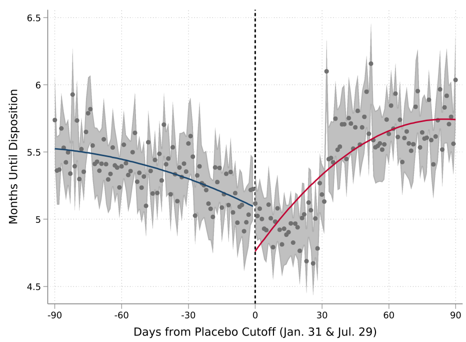
	(1)	(2)	(3)	(4)
Filed After Cutoff	-0.205	0.053	0.012	-0.258**
	(0.149)	(0.118)	(0.134)	(0.107)
Cutoff Dates	2 months left	1 month left	1 month right	2 months right
Dep. Var. mean	5.4	5.39	5.4	5.4
Polynomial order	First	First	First	First
Controls	✓	✓	✓	✓
Bandwidth (days)	22.6	18.6	19.3	20.6
Effective N	49,858	43,197	46,214	49,434
N (left)	25,456	20,321	22,346	23,488
N (right)	24,402	22,876	23,868	25,946

Note: Table presents placebo RD estimates of the effect of reporting time on months until motion disposition using counterfactual six-month publication and cutoff dates (shifted either one or two months to the left or right of actual dates). All columns are estimated with local linear polynomials using triangular kernels.

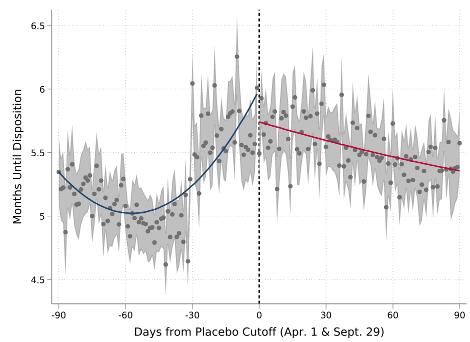
Table A7—RD Estimates for Motion & Appellate Outcomes, Alternative Bandwidths

BW (days):	7	14	21	28	35	42
% Granted	-0.01	-0.02	-0.02	-0.02	-0.01	-0.01
	(0.02)	(0.02)	(0.01)	(0.01)	(0.01)	(0.01)
% Denied	0.01	0.01	0.01	0.01	0.01	0.01
	(0.03)	(0.02)	(0.02)	(0.01)	(0.01)	(0.01)
% Appealed	0.02	0.00	-0.00	-0.01	-0.01	-0.01
	(0.02)	(0.02)	(0.01)	(0.01)	(0.01)	(0.01)
% Affirmed	-0.00	0.03	0.04**	0.04**	0.03**	0.03**
	(0.02)	(0.02)	(0.02)	(0.02)	(0.02)	(0.01)

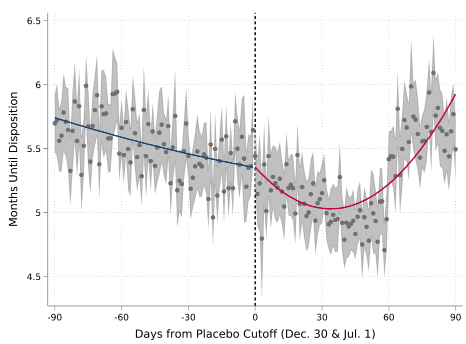
Note: Table presents RD estimates of the effect of reporting time on motion and appellate outcomes using a variety alternative bandwidths. All columns are estimated with first first-order local polynomials and triangular kernels. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.



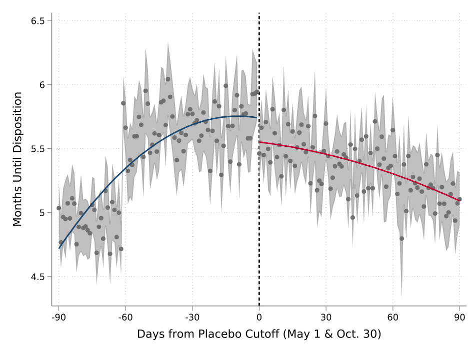
(a) One Month Left



(b) One Month Right



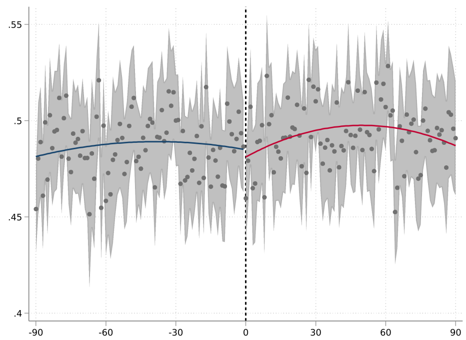
(c) Two Months Left



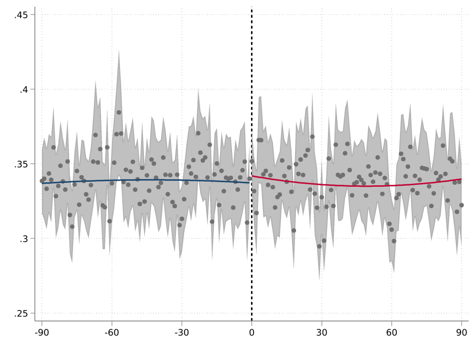
(d) Two Months Right

FIGURE A3. RD ESTIMATES OF EFFECT ON MONTHS UNTIL DISPOSITION, PLACEBO CUTOFFS

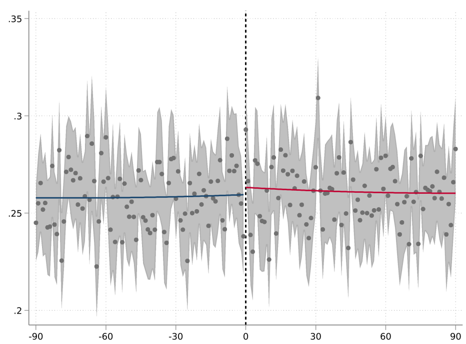
Note: Figures show RD placebo plots of the effect of six months of additional reporting time on months until motion disposition using counterfactual six-month publication and cutoff dates. Publication and cutoff dates are shifted either one or two months to the left or right. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). All columns are estimated with local linear polynomials using triangular kernels.



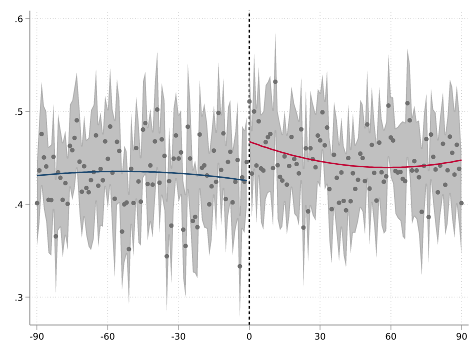
(a) % Granted



(b) % Denied



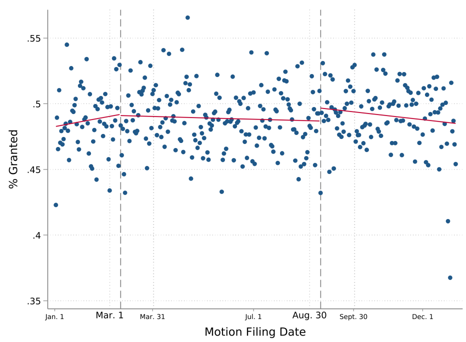
(c) % Appealed



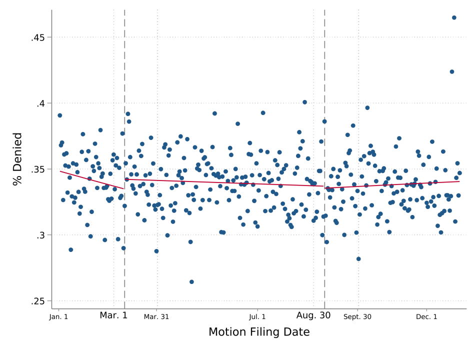
(d) % Affirmed on Appeal

Figure A4. RD Plots: Motion and Appellate Outcomes

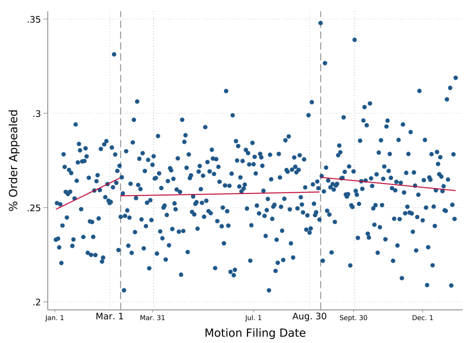
Note: Figures plot average motion and appellate outcomes against the motion's filing date relative to the two six-month list cutoff dates of March 1 and August 30. The sample comprises all summary judgment motions filed from 2005–2014 with known dispositions.



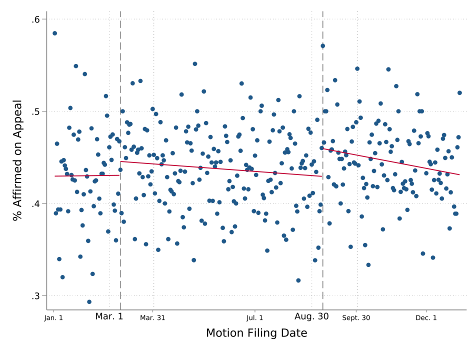
(a) % Granted



(b) % Denied



(c) % Appealed



(d) % Affirmed on Appeal

FIGURE A5. SCATTERPLOTS: MOTION AND APPELLATE OUTCOMES

Note: Figures plot average motion and appellate outcomes against the motion's calendar filing date. Best-fit lines are separately estimated for the periods January 1st–February 29th, March 1–August 29th, and August 30–December 31st. The sample comprises all summary judgment motions filed from 2005–2014 with known dispositions.

Table A8—Donut RD Estimates for Motion & Appellate Outcomes

Donut radius (days):	3	6	9	12	15	18
% Granted	-0.01	-0.01	-0.01	-0.01	-0.01	-0.01
	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)
% Denied	0.01	0.01	0.01	0.01	0.01	0.01
	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)
% Appealed	-0.01	-0.01	-0.01	-0.01	-0.00	-0.00
	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)
% Affirmed	0.04**	0.03**	0.03**	0.03**	0.03**	0.03**
	(0.02)	(0.02)	(0.02)	(0.01)	(0.01)	(0.01)

Note: Table presents "Donut" RD estimates of the effect of reporting time on motion and appellate outcomes using a variety of donut radii. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). All columns are estimated with triangular kernels using first-order local polynomials. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Table A9—Placebo Tests: RD Estimates for Motion & Appellate Outcomes, Counterfactual Cutoffs

Cutoff Dates:	2 Mos. Left	1 Mo. Left	1 Mo. Right	2 Mos. Right
% Granted	-0.01	-0.01	-0.01*	-0.00
	(0.01)	(0.01)	(0.01)	(0.01)
% Denied	0.01	0.01	0.00	0.00
	(0.01)	(0.01)	(0.01)	(0.01)
% Appealed	-0.01	-0.01	0.01	-0.01
	(0.01)	(0.01)	(0.01)	(0.01)
% Affirmed	0.01	-0.02	0.01	-0.02
	(0.02)	(0.02)	(0.02)	(0.01)

Note: Table presents placebo RD estimates of the effect of reporting time on motion and appellate outcomes using counterfactual six-month publication and cutoff dates. Publication and cutoff dates are shifted either one or two months to the left or right. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). All columns are estimated with local linear polynomials using triangular kernels.

Table A10—RD Estimates: Effect of six months additional reporting time on motion & appellate outcomes, clustered at Julian date of filing

	(1)	(2)	(3)	(4)
% Granted	-0.02	-0.01	-0.02	-0.02
	(0.01)	(0.01)	(0.02)	(0.02)
% Denied	0.01	0.01	0.01	0.02
	(0.01)	(0.01)	(0.02)	(0.02)
% Appealed	0.00	-0.01	0.02	-0.00
	(0.01)	(0.01)	(0.02)	(0.02)
% Affirmed	0.05**	0.04**	0.06**	0.04**
	(0.02)	(0.02)	(0.03)	(0.02)
Polynomial order	First	First	Second	Second
Controls		✓		✓

Note: Table presents bias-corrected RD estimates of the effect of six months additional reporting time on motion & appellate outcomes. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). Effects are estimated with first or second order local polynomials using triangular kernels. Columns indicating controls include list (i.e., March or September) fixed effects, a dummy for the part (plaintiff or defendant) filing the motion, and nature-of-suit and filing-year fixed effects. Robust standard errors clustered at the Julian date of filing are in parentheses.

Table A11—RD Estimates: Effect of six months additional reporting time on motion & appellate outcomes, clustered at judge-level

	(1)	(2)	(3)	(4)
% Granted	-0.02	-0.01	-0.03	-0.03
	(0.01)	(0.01)	(0.02)	(0.02)
% Denied	0.01	0.01	0.01	0.02
	(0.01)	(0.01)	(0.02)	(0.02)
% Appealed	0.00	-0.01	0.02	0.00
	(0.01)	(0.01)	(0.02)	(0.02)
% Affirmed	0.04**	0.04*	0.07*	0.04
	(0.02)	(0.02)	(0.04)	(0.03)
Polynomial order	First	First	Second	Second
Controls		✓		✓

Note: Table presents bias-corrected RD estimates of the effect of six months additional reporting time on motion & appellate outcomes. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). Effects are estimated with first or second order local polynomials using triangular kernels. Columns indicating controls include list (i.e., March or September) fixed effects, a dummy for the part (plaintiff or defendant) filing the motion, and nature-of-suit and filing-year fixed effects. Robust standard errors clustered at the judge-level are in parentheses.

Table A12—OLS Estimates: Effect of one month additional reporting time on months until motion disposition, motion disposition dates imputed from case closures

	(1)	(2)	(3)	(4)
Reporting Time (Months)	0.093***	0.106***	0.099***	0.101***
	(0.011)	(0.012)	(0.011)	(0.010)
Observations	362,117	362,117	362,114	362,114
Dep. Var. mean	✓	✓	✓	✓
Case & Motion Controls		✓	✓	✓
Calendar Trends			✓	✓
DistrictYear FEs				✓
Dep. Var. mean	5.79	5.79	5.79	5.79
Indep. Var. mean	10.05	10.05	10.05	10.05

Note: Table presents OLS estimates of the effect of additional reporting time on months until motion disposition. For motions with unknown dispositions but known case closure dates, the case closure date is used as the motion disposition date. MSE-optimal bandwidths are selected following the approach of Calonico, Cattaneo and Titiunik (2014). Effects are estimated with first or second order local polynomials using triangular kernels.

Table A13—RD Estimates: Effect of six months additional reporting time on months until motion disposition, motion disposition dates imputed from case closures

	(1)	(2)	(3)	(4)
Filed After Cutoff	0.819	0.337	1.233*	0.458
	(0.535)	(0.245)	(0.670)	(0.300)
Dep. Var. mean	5.79	5.79	5.79	5.79
Polynomial order	First	First	Second	Second
Controls		✓		✓
Bandwidth (days)	24.5	21.6	30.7	35.5
Effective N	97,704	86,441	123,254	144,333
N (left)	48,012	42,137	61,157	71,016
N (right)	49,692	44,304	62,097	73,317

Note: Table presents RD estimates of the effect of additional reporting time on months until motion disposition. For motions with unknown dispositions but known case closure dates, the case closure date is used as the motion disposition date.

Supplemental Appendix: Bunching Methodology & Additional Results

Figure 1 shows clear bunching of motion dispositions at the six-month list deadlines. To quantify the aggregate effect of the six-month list on motion-processing time, I implement a notched bunching estimator adapted from Kleven and Waseem (2013) and Kleven (2016).

Let d be the number of days from motion filing to motion disposition, and let l be the number of days from the filing date to the six-month list deadline (i.e., the motion's reporting time in days). I observe data generated by the density functions $f_1^l(d)$—the actual distribution of motion resolution times given a six-month reporting time of l—for $l \in L$ and the goal is to obtain a counterfactual distribution $f_0(d)$. My setting and estimator are similar to those of Gruber, Hoe and Stoye (2023), who study the effect of a policy in the U.K. that penalized emergency room providers for failing to admit patients within four hours of their arrival. Unlike their setting, however, where doctors were subjected to a single universally-applicable wait time target, my setting involves many *different* notches corresponding to the amount of reporting time a motion receives, between approximately seven and thirteen months.

I first group motion dispositions into five-day bins indexed by j—for example, motions decided in five days or less, motions decided in six to ten days, and so on. From there, I can bin estimate separate counterfactual distributions $\hat{f}_0^l(d)$ for each possible reporting time l using a series of polynomial regressions of the form:

$$\text{(B1)} \qquad c_j^l = \sum_{i=0}^{p} \beta_i^l \cdot (d_j)^i + \sum_{i=d_-}^{d_+} \gamma_i^l \cdot \mathbb{1}[d_j = i] + \sum_{r \in R} \rho_r^l \cdot \mathbb{1}\left[\frac{d_j}{r} \in \mathbb{N}\right] + \nu_j^l,$$

where c_j^l is the number of individual motions with reporting time l disposed of in bin j, d_j is the maximum duration of a motion disposed of in bin j (for example, $d_j = 5$ for motions adjudicated in 1–5, $d_j = 10$ for motions decided in 6–10 days, and so on), and p is the order of the polynomial ($p = 8$ in my baseline specification). The term $\sum_{r \in R} \rho_r \cdot \mathbb{1}\left[\frac{d_j}{r} \in \mathbb{N}\right]$ reflects the inclusion of round-number fixed effects. These are necessary to account for the fact that dispositions tend to spike on multiples of seven, 30, and 365.

The range $[d_-, d_+]$ is the "excluded window" of excess and missing mass around the notch point l. In addition to smoothness of the counterfactual distribution, the key assumption of the bunching estimator is that bunching responses are entirely confined to the excluded window. For example, I assume that judges might respond to the March 31 deadline by postponing dispositions that otherwise would have occurred in late February or early March, and they might similarly respond by expediting dispositions that otherwise would have occurred in April or May; they are unlikely, however, to either postpone a disposition they would have otherwise reached much sooner or expedite a disposition they would have otherwise reached much later.

This is a relatively strong assumption, and it is likely most defensible for motions filed on or just prior to March 1 and August 30 (i.e. those with maximum reporting time). Specifically, when judges have a full 13 months to review a motion before it becomes eligible for the six-month list, the pull of the six-month list is at its weakest. It is reasonable to assume that, during the first several months of a motion's pendency, judicial behavior will not be greatly affected by a deadline that is still many months away. Somewhat more intuitively, it is reasonable to believe that there is little difference between a distant deadline that hardly ever binds and a regime with no deadline at all.

The estimated counterfactual distribution is defined as the predicted values of equation (B1) omitting the contribution of the excluded window dummies so that $\hat{c}_j^l = \sum_{i=0}^{p} \hat{\beta}_i \cdot (d_j)^i + \sum_{r \in R} \rho_r^l \cdot \mathbb{1}\left[\frac{d_j}{r} \in \mathbb{N}\right]$. I depart from the Kleven and Waseem (2013)—which chooses the lower bound of the exclusion window d_- by visual inspection—by implementing a Quandt Likelihood Ratio (QLR) test. The QLR test is frequently used to identify structural breaks in time-series data. For my purposes, the QLR test identifies the duration bin d_- where dispositions have most strongly broken from their prior trend. The upper bound d_+ chosen recursively by starting at an initial value $d_+^0 \approx l$, estimating equation (B1), and increasing d_+ by small increments until we identify the value of d_+ that minimizes the difference between estimated excess mass $\hat{B}$ in the range $[d_-, l]$ and estimated missing mass $\hat{M}$ in the range $[l, d_+]$.

While my data make it feasible to estimate these separate reporting time-specific counterfactual distributions, I instead choose to estimate a single counterfactual distribution $\hat{f}_0(d)$ using only motions with close to the maximum possible reporting time (i.e., 390 days or greater). I make this simplification for two reasons: first, it makes sense to impose a single density function since, theoretically, the counterfactual distribution should be the same for all reporting times; and second, because far-off deadlines are similar to no deadline at all, the core bunching assumption (i.e., that the distribution of adjudications is unaffected by the six-month list outside of a local "exclusion window" around the deadline itself) is most likely to be satisfied in the maximum reporting time scenario. Figure B1 plots the distribution of motion resolution times for motions with reporting times of 390 days or greater, along with the fitted counterfactual distribution.[15] Using the approach described above, I estimate an exclusion window beginning 15 days to the left of the six-month list deadline (375 days) and ending 95 days to the right (485 days).

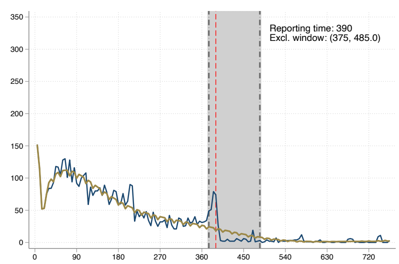


Figure B1. Actual vs. Counterfactual Dispositions, Maximum Reporting Time

Note: Figure plots the observed and counterfactual frequencies of motion disposition times (in days) for motions with maximum reporting time (at least 390 days). Counterfactual distribution is estimated using an 8th-order polynomial regression. Exclusion window is shaded in gray.

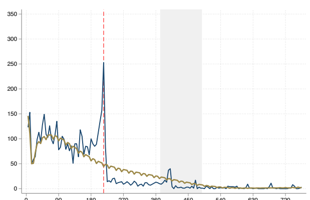
With a counterfactual distribution in hand, I can then estimate the aggregate effect of the six-month list by comparing the actual and counter-

[15]Notably, the observed density of motion dispositions appear to spike slighlty at approximately 215 days, which aligns with the six-month list deadlines for more exposed motions. This suggests that some judges may be responding to the six-month list even when they have a full 13 months to resolve a motion. This is not entirely surprising since, as discussed above, the precise six-month list inclusion criteria are somewhat opaque and may not always be known even by the judges themselves. To the extent that judges are responding to deadlines that are not actually binding on them, this would tend to bias my estimates of the effect of the six-month list downward.

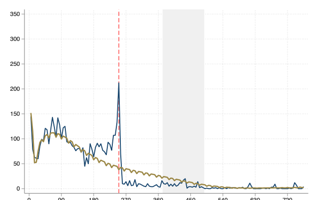
factual distributions. Specifically, I calculate the ratio $T/\hat{T}$, where $T = \sum_{j\in J}\sum_{l\in L} c_j^l \cdot d_j$ represents the total actual disposition time for all summary judgment motions in my sample, and $\hat{T} = \sum_{j\in J}\sum_{l\in L} \hat{c}_j^l \cdot d_j$ represents the estimated total counterfactual disposition time for the same sample of motions. I assume that the counterfactual for any given reporting time l is equal to the counterfactual estimated from maximum reporting time data ($l = 390$ days) scaled by the number of filings with reporting time l, so that $\hat{c}_j^l = \hat{c}_j^3 90 * \frac{\sum_j c_j^l}{\sum_j c_j^3 90}$.

Table B1 presents ratios $T/\hat{T}$ for a variety of different polynomial orders and exclusion window lower bounds. A ratio of less than one suggests that the six-month list has reduced aggregate disposition times relative to the scenario without a six- month list, and a value of greater than one would suggest that the six-month list has actually slowed aggregate dispositions. Across all specifications, I estimate that the six-month list reduces total motion disposition time by approximately 3.5–4.5%.

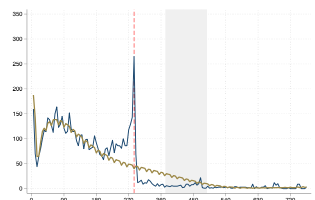
As discussed above, our setting is complicated by the fact that motions are subject to a variety of different reporting times—more than 180 in total—depending upon the day of the year on which they were filed. Thus, we might hypothesize that the extent of bunching is likely to vary with the amount of reporting time. Figure B2 confirms our intuition by plotting the actual versus counterfactual density of motion disposition times for six selected reporting time bins, ranging from the minimum (just over seven months, depicted in the northwest sector of the plot) to the maximum (just over thirteen months, depicted in the southeast sector of the plot). We can see that the bunching is prominent for all possible reporting times, but it is most prominent for motions with a relatively low amount of reporting time.



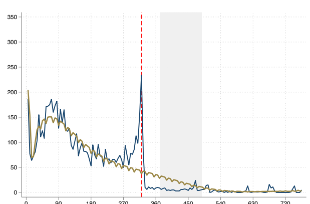
(a) 215 days R.T.



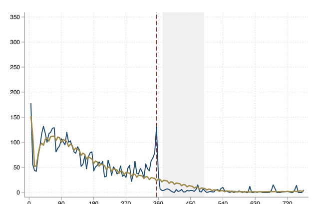
(b) 250 days R.T.



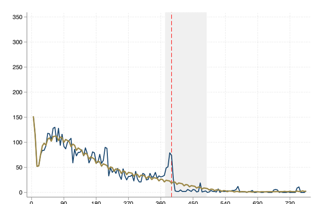
(c) 285 days R.T.



(d) 320 days R.T.



(e) 355 days R.T.

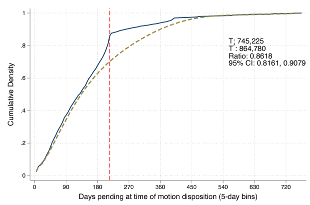


(f) 390 days R.T.

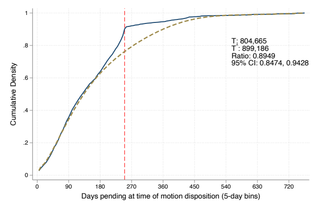
Figure B2. Counterfactual & Observed Disposition Times (Days) for Six Selected Reporting Times

Note: Figure plots observed and counterfactual disposition times (in days) for six different reporting times. Counterfactual distributions are estimated from motions with maximum reporting time ($\geq$ 390 days) using an 8th-order polynomial regression. Exclusion windows are shaded in light gray.

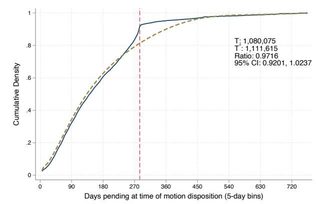
Finally, Figure B3 plots observed versus counterfactual CDFs for the same six reporting time bins. Reporting time-specific ratios $T^l/\hat{T}^l$ and confidence intervals are indicated in the northeast corner of each plot.[16]



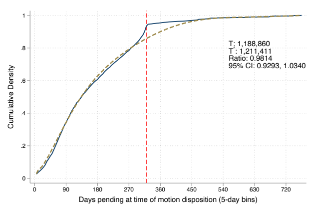
(a) 215 days R.T.



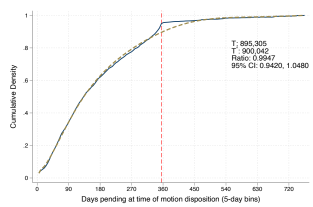
(b) 250 days R.T.



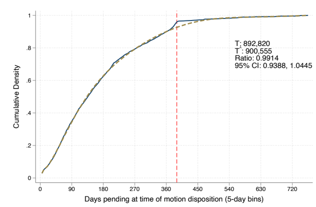
(c) 285 days R.T.



(d) 320 days R.T.



(e) 355 days R.T.



(f) 390 days R.T.

FIGURE B3. CDFS OF COUNTERFACTUAL & OBSERVED DISPOSITION TIME (DAYS) FOR SIX SELECTED REPORTING TIMES

Note: Figure plots observed and counterfactual disposition time CDFs (in days) for six different reporting times. Counterfactual distributions are estimated from motions with maximum reporting time ($\geq$ 390 days) using an 8th-order polynomial regression. Exclusion windows are shaded in light gray.

[16]Standard errors and confidence intervals are obtained with a Wild bootstrap procedure.

Table B1—Robustness Checks: Bunching Estimates by Polynomial Order and Lower Bound

	Lower bound of exclusion window			
Order (p)	**360**	**370**	**375**	**380**
6	0.9605	0.9581	0.9566	0.9609
	(0.8916–1.0001)	(0.8901–0.9995)	(0.8881–0.9957)	(0.8796–1.0056)
7	0.9573	0.9546	0.9588	0.9637
	(0.9065–1.0077)	(0.9056–0.9966)	(0.9006–1.0021)	(0.9006–1.0094)
8	0.9532	0.9548	0.9584	0.9590
	(0.9003–1.0176)	(0.9072–1.0036)	(0.9076–1.0098)	(0.9032–1.0076)
9	0.9453	0.9586	0.9643	0.9603
	(0.8987–0.9982)	(0.8979–1.0093)	(0.8953–1.0026)	(0.8959–1.0041)

Note: Table presents estimates of the ratio $T/\hat{T}$ for varying polynomial orders and lower bounds of the exclusion window. Polynomial of order 8 and lower bound of 375 are the baseline. 95% confidence intervals are presented in parentheses below each estimate.